

Tests of stochastic dominance for multivariate distributions

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- ▶ In multivariate settings comparing the marginal distributions of A and B for every variable may not be appropriate to ascertain stochastic dominance (and partial orderings)
- ▶ Instead it may be necessary to perform tests on functions of the joint distribution functions!
- ▶ Whether we need to check the joint distributions or not depends on the cross derivatives of the class of welfare functions for which we are trying to establish partial orderings
- ▶ In other words, it depends on whether the class of functions is characterized by neutrality among the dimensions or not (i.e. they are either substitutes or complements, or both at times)

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 - ▶ Declare $A = B$ if and only if $A = B$ for all variables (x_1, \dots, x_V)
 - ▶ Declare indeterminacy (i.e. no dominance and no homogeneity) if and only if $\exists_{x_i} | ADB$ and $\exists_{x_{j \neq i}} | BDA$

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We are going to study two of these for poverty functions: one for continuous variables and one for a combination of continuous and discrete variables (both by Duclos, Sahn and Younger, 2006)

Refreshing dominance conditions and further relevance in multivariate settings

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A stochastic dominance condition is a relationship between functions of the distribution functions of a pair of samples, A and B, which in turn establishes a consistent ranking of A versus B across a range of evaluation functions belong to a specific class. This is the form of a typical stochastic dominance condition (the left hand side is by definition the relationship between the functions):

$$R[G_A, G_B] \iff W_A(x) \geq W_B(x) \forall x \wedge \exists x | W_A(x) > W_B(x) \forall W \in \mathcal{W}^D$$

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- ▶ It is a fantastic result.
- ▶ It allows for a wide range of conditions.
- ▶ The conditions provided allow for poverty comparisons robust not just to dimension weights but chiefly to different poverty lines representing different identification criteria (e.g. union, intersection, intermediate)
- ▶ Main drawback: it only considers classes of welfare functions characterized by substitutability between dimensions

The test of Duclos et al. (2006): preliminary notation

The conditions work over additive poverty functions:

$$P(\lambda) = \int \int_{\Lambda(\lambda)} \pi(x, y; \lambda) dF(x, y)$$

The individual poverty function is characterized by:

$\pi(x, y; \lambda) \geq 0$ if $\lambda(x, y) \leq 0$ and $\pi(x, y; \lambda) = 0$ otherwise

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The multiplicative FGT surface (upon which the conditions are derived) is:

$$P^{\alpha_x, \alpha_y}(z_x, z_y) = \int_0^{z_y} \int_0^{z_x} (z_x - x)^{\alpha_x} (z_y - y)^{\alpha_y} dF(x, y)$$

where α_x and α_y are positive integers

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The test of Duclos et al. (2006): dominance conditions for the examples

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In their paper Duclos et al. (2006) show, for their examples, that:

$$\Delta P(\lambda) > 0 \forall P(\lambda) \in \prod^{1,1}(\lambda^*) \iff \Delta P^{0,0}(x, y) > 0 \forall x, y \in \Lambda(\lambda^*)$$

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These conditions can be extended into higher orders of dominance by dimension (the α 's) and more dimensions

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3. For instance we can test $H_0 : P_A^{\alpha_x, \alpha_y} \leq P_B^{\alpha_x, \alpha_y}$ against $H_a : P_A^{\alpha_x, \alpha_y} > P_B^{\alpha_x, \alpha_y}$

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4. Rejection of that null hypothesis in favor of that alternative would require that $Z_{\Delta P} \geq Z_{critical} \forall x, y$ and $\exists x, y | Z_{\Delta P} > Z_{critical}$

Testing strategy: the covariance matrices

As in the univariate case we have the case of independent samples (e.g. two countries) and dependent samples (e.g. panel data)

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- ▶ Now we need to account for more dimensions
- ▶ We are estimating the standard errors of differences of FGT as opposed to differences in functions of probabilities (e.g. the factorial is not considered)

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$$\begin{aligned} & \text{cov}[\widehat{P}_{A,t}^{\alpha_x, \alpha_y}(z_x, z_y), \widehat{P}_{A,t+1}^{\alpha_x, \alpha_y}(z_x, z_y)] = \\ & \left(\frac{1}{N}\right)^2 \sum_{i=1}^N (z_y - y_{i,t})_+^{\alpha_y} (z_x - x_{i,t})_+^{\alpha_x} (z_y - y_{i,t+1})_+^{\alpha_y} (z_x - x_{i,t+1})_+^{\alpha_x} \\ & - \left(\frac{1}{N}\right) \frac{1}{N} \sum_{i=1}^N (z_y - y_{i,t})_+^{\alpha_y} (z_x - x_{i,t})_+^{\alpha_x} \frac{1}{N} \sum_{i=1}^N (z_y - y_{i,t+1})_+^{\alpha_y} (z_x - x_{i,t+1})_+^{\alpha_x} \end{aligned}$$

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- ▶ For the dominance conditions we use the following conditional FGT function: $P^\alpha(k, k^*; z) = \int_0^z (z - x)^\alpha f(x|k, k^*) dx$

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- ▶ For the dominance conditions we use the following conditional FGT function: $P^\alpha(k, k^*; z) = \int_0^z (z - x)^\alpha f(x|k, k^*) dx$
- ▶ The generic aggregate poverty index is defined over several poverty lines: $P(z(1, 1); z(1, 2); \dots; z(1, K^*); \dots; z(K, K^*))$

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It assumes ordinality in all discrete variables (sensible)

The test of Duclos et al. (2006) for one continuous and several discrete variables: the example they provide

They find a dominance condition for a general class named $\dot{\Pi}^1(z(1, 1); z(1, 2); \dots; z(1, K^*); \dots; z(K, K^*))$

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4. $\pi_{k, k^*}^{(1)} \leq \pi_{k, k^*+1}^{(1)} \leq 0 \quad \forall x, k, k^*$
5. $\pi_{k, k^*}(z(k, k^*)) = 0 \quad \forall k = 1, \dots, K; k^* = 1, \dots, K^*$

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$\dot{\Pi}^1(z(1, 1); z(1, 2); \dots; z(1, K^*); \dots; z(K, K^*))$ is:

$$\Delta P(\zeta(1, 1), \dots, \zeta(K, K^*)) > 0 \forall P(\zeta(1, 1), \dots, \zeta(K, K^*)) \in \dot{\Pi}^1$$

$$\wedge \forall \zeta(k, k^*) \in [0, z(k, k^*)], k = 1, \dots, K; k^* = 1, \dots, K^*$$

$$\longleftrightarrow \sum_{k=1}^i \sum_{k^*=1}^j \Delta P^0(k, k^*; \zeta) > 0, \forall \zeta \in [0, z(i, j)]$$

$$\wedge i = 1, \dots, K; j = 2, \dots, K^*$$

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The procedures are very similar as in the previous cases, for instance:

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$$\sum_{k=1}^i \sum_{k^*=1}^j \widehat{P}^\alpha(k, k^*; \zeta) = \frac{1}{N} \sum_{m=1}^N (\zeta(k_m, k^*_m) - x_m)_+^\alpha I(k_m \leq i \wedge k^*_m \leq j)$$

The test of Duclos et al. (2006) for one continuous and several discrete variables: estimations of statistics

Another example, the covariance of dependent samples with one continuous variable and one discrete variable:

$$\begin{aligned} \text{cov}\left(\sum_{k=1}^j \widehat{P}_t^\alpha(k; \zeta), \sum_{k=1}^j \widehat{P}_{t+1}^\alpha(k; \zeta)\right) = \\ \frac{1}{N^2} \sum_{m=1}^N (\zeta(k_{m,t}) - x_{m,t})_+^\alpha (\zeta(k_{m,t+1}) - x_{m,t+1})_+^\alpha I(k_{m,t} \leq j \wedge k_{m,t+1} \leq j) \\ - \left(\frac{1}{N}\right) \sum_{k=1}^j \widehat{P}_t^\alpha(k; \zeta) \sum_{k=1}^j \widehat{P}_{t+1}^\alpha(k; \zeta) \end{aligned}$$