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Evaluating the Effects of Housing Interventions on Multidimensional Poverty: The Case of TECHO-Argentina

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Abstract

The objective of this paper is to evaluate the effect of the NGO TECHO's emergency housing programme on multidimensional poverty. It employs a quasi-experimental 'pipeline' evaluation design and is based on household survey data from 34 informal settlements in Buenos Aires, Argentina. The aim is to demonstrate the additional insights that can be gained from using a multidimensional framework based on the Alkire and Foster (2011) method to evaluate a programme's impact. The results indicate that the programme reduces both the incidence and the intensity of poverty and causes the multidimensional poverty measure to fall by more than half. The magnitude of the effect is greater for the households that initially were the poorest. Privacy, interpersonal relations and psychological health are the dimensions that contribute the most to explaining the decline in multidimensional deprivation.

Keywords: Multidimensional poverty, impact evaluation, housing, informal settlements.

JEL classification: C13, I32, R21

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1. Introduction

The right to adequate housing is a basic human right.¹ The United Nations asserts that for shelter to be adequate it must satisfy the following conditions: 1) security of tenure; 2) availability of basic services and infrastructure; 3) affordability; 4) habitability; 5) accessibility; 6) located in proximity to schools, work and other facilities and 7) respect the expression of cultural identity (UNOHCHR, 2009). The condition of habitability requires housing to guarantee physical safety, adequate space, and protection from cold, heat, rain, wind and other threats to health and structural hazards (UNOHCHR, 2009). The right to housing is also interrelated with other human rights such as the right to work, health and education.

The New Urban Agenda adopted in 2016 in Habitat III² stresses the need for governments to put ‘housing at the centre’ of urban development planning. This new focus represents an important shift from the ‘enabling shelter strategy’ introduced in 1988 and later elaborated on in the Istanbul Declaration and Habitat Agenda, which proposed a market-led approach in which governments facilitate or motivate the deployment of resources for housing production by all actors—public, private, civil society and the people themselves (UN-Habitat, 2016). The change in focus was motivated by the recognition that market-led policies have failed to provide adequate, affordable housing to low-income families (UN-Habitat, 2016) and to reduce the number of people living in urban slums (UNDP, 2015). Target 11.1 of the Sustainable Development Goals—to ensure access for all to adequate, safe and affordable housing and basic services, and upgrade slums by 2030—is in line with this new commitment to improving living conditions in urban slums (United Nations Sustainable Development Solutions Network, 2018).

While the percentage of the global urban population living in slums declined from 46% to 30% between 1990 and 2014, the absolute number increased to 880 million, compared with 689 million in 1990 (UN-Habitat, 2016:14). The estimated urban slum population is 563 million in Asia, 212 million in Africa and 104 million in Latin America (UN-Habitat, 2016:58). Urban slums are characterized by insecure tenancy, inadequate provision of public services, and poor quality and overcrowded dwellings. Moreover, 980 million urban households live in an inadequate dwelling (UN-Habitat, 2016).

In the context of an inadequate supply of affordable urban housing and mortgage credit, low-income families construct their own housing on unoccupied lands (Rojas, 2016). According to some accounts,

¹ The 1948 UN Declaration of Human Rights recognized adequate housing as part of the right to an adequate standard of living. This right was confirmed in the 1966 International Covenant on Economic Social and Cultural Rights.

² Habitat III was the United Nation’s Conference on Housing and Sustainable Urban Development which took place in Quito, Ecuador, October 17–20, 2016. [Link](#).

more than half of all housing in the Global South can be classified as socially produced, that is, constructed by the future residents alone or with the help of non-profit organizations (Ortiz Flores, 2011).

Civil society is a central component of this collective and organized form of urban development. Social movements, cooperatives and self-help organizations coordinate land occupations, collectively produce housing, introduce neighbourhood improvements and provide public service connections (Ortiz, 2011; Wagner, 2011). Non-governmental organizations support these grassroots efforts through programmes to upgrade housing, facilitate credit, broaden access to public services and advocate for public sector actions to improve habitat. TECHO is one of the NGOs working to improve living conditions in urban slums in Latin America that has the broadest geographical scope.

The objective of the paper is to evaluate the effect of TECHO's emergency housing programme on multidimensional poverty. It employs a quasi-experimental 'pipeline' evaluation design and is based on household survey data collected in 34 informal settlements in Buenos Aires, Argentina. The paper seeks to broaden the literature on the impact of TECHO's housing programme (Simonelli et al., 2013; Galiani et al., 2016; Mitchell, Macció and Mariño Fages, 2016) by measuring its effects on multidimensional poverty using the method proposed by Alkire and Foster (2011).

There is an expanding body of literature on the effectiveness of providing emergency housing on-site in informal settlements. The evaluation carried out by Galiani et al. (2016) in Mexico, El Salvador and Uruguay, which used an experimental evaluation design, found that a TECHO house produced an increase in subjective measures of quality of life and satisfaction with housing and a reduction in the incidence of diarrhoea in children (in Mexico and El Salvador), but had no effect on the possession of durable goods, complementary housing investments, the incidence of robbery, fertility, income or labour market outcomes. Simonelli et al. (2013) carried out a smaller scale evaluation of TECHO's programme in Argentina, which focussed on the effects of emergency housing on sleep. Based on a before-after methodology, the authors concluded that the programme was associated with reductions in levels of stress and dissatisfaction with housing and improvements in sleep quality and perception of quality of life. Mitchell, Macció and Mariño Fages (2016) show that the TECHO programme in Buenos Aires had a positive effect on privacy, security, interpersonal relations, psychological health and perception of quality of life, and also increased the percentage of families with school-age children that have a tranquil place to study. Another relevant evaluation of emergency housing is Loschman et al. (2015), which studied the effects of emergency shelters provided to refugees returning to Afghanistan.

The contribution of this paper is to demonstrate that when evaluating interventions that affect multiple wellbeing dimensions, additional insights can be gained from applying a multidimensional framework instead of measuring the effects on individual outcome variables. First, it enables one to obtain a summary

measure of the overall effect of the programme on deprivation. Consolidated outcome measures are increasingly used in the impact evaluation literature as a means for resolving the increased probability of type 1 error (a false positive finding) when carrying out multiple hypothesis tests (Anderson, 2008). The Alkire and Foster (2011) method provides a framework for constructing a summary outcome measure based on normative criteria as opposed to purely empirical methods. Second, the method enables one to analyse programme effects on the incidence of multidimensional poverty, the intensity of deprivation experienced by the multidimensionally poor and on a distribution-sensitive measure of poverty incidence. This allows one to analyse the extent to which the policy enables households to escape multidimensional poverty or reduces the intensity of deprivation for those that continue to be multidimensionally poor. Third, the method offers a method for examining the heterogeneity of the effects on households with different levels of initial deprivation. Fourth, as the Alkire-Foster measures are decomposable by indicator, this property can be used to analyse which types of deprivation contribute most to poverty before and after treatment.

The paper is structured as follows. Section two describes TECHO's emergency housing programme. Section three explains the Alkire-Foster method. Section four outlines the evaluation design, data sources and the empirical strategy used to evaluate the programme's effects on multidimensional poverty. Section four presents the empirical results. The paper concludes with a summary of the main findings and discusses the benefits and limitations of the analysis.

2. TECHO's emergency housing programme

TECHO, created in 1997 by a group of young people in Chile, today works in 19 Latin American and Caribbean countries. The organization's stated goal is to fight poverty by working with the residents of informal settlements to improve habitat and foster community development. Using a participatory process, TECHO's small staff and abundant supply of volunteers work with communities to identify needs and then develop programmes that respond to those needs, including emergency housing, education, employment, youth and microfinance programmes.

TECHO's emergency housing programme provides a basic 18-square-meter dwelling to families that live in informal settlements and experience a severe deficit in the size and quality of housing. The emergency house is made of prefabricated wood panels, has an insulated tin roof and rests on 15 pilings that elevate the dwelling from ground humidity and flooding. The programme does not provide water, sanitation or electrical connections. The cost of materials for each unit is approximately US\$1,000.

Groups of four to eight volunteers construct the emergency dwelling, together with the future residents, on the plot of land where the household resides. This enables the family to continue living in the same neighbourhood and avoids the severing of social and community ties, but also means that when the size of the land parcel is small, the construction of the new dwelling can exacerbate overcrowding. In some cases, the original dwelling must be dismantled in order to make space for the new construction. As the household remains in the same neighbourhood, the programme does not address or improve any of the other common problems in informal settlements, such as insecure tenancy, lack of basic infrastructure or environmental hazards. This characteristic of the TECHO programme appears to be unique. In fact, we have not identified any other NGO programme that provides a new dwelling but does not facilitate access to water and sanitation.

TECHO-Argentina's participant selection process consists of first conducting, in each neighbourhood, a detection survey that collects information on the sociodemographic characteristics of the members, economic situation of the household, access to public services and characteristics of the dwelling. This information is then used to classify households according to the level of need. Households considered to have high and medium-high need are eligible to participate in the programme. Before construction the households must sign a contract in which they agree to pay around 10% of the cost of the materials, attend preconstruction neighbourhood meetings and clear and prepare the plot of land.

3. Alkire-Foster method

The Alkire-Foster method provides a framework for measuring poverty based on a counting approach that assesses the number of deprivations that people experience in multiple dimensions of life. It is composed of identification and aggregation stages.

In the identification stage, it is first necessary to select the dimensions of analysis and one or more indicators within each dimension. For each indicator one must choose a deprivation cut-off, which expresses the minimum level of achievement necessary to be non-deprived, and a relative weight, such that the sum of the indicator weights equals one. These criteria are used to calculate for each household (or person) the weighted share of deprivations, or deprivation score c_i . Then it is necessary to set a second threshold, the poverty cut-off or k , which is the proportion of weighted deprivations that a household must experience in order to be considered multidimensionally poor. If a household has a deprivation score greater than or equal to the poverty cut-off ($c_i \geq k$), it is considered to be multidimensionally poor. Once the identification stage is complete, one can define a deprivation score vector, $c = [c_i]$, in which c_i is the deprivation score of individual i . Additionally, one can construct a censored deprivation score vector $c(k) = [c_i(k)]$, in which $c_i(k) = c_i$ if the person is multidimensionally poor ($c_i \geq k$) and $c_i(k) = 0$ otherwise.

The aggregation stage of the Alkire-Foster method is based on the extension of the Foster-Greer-Thorbecke family of unidimensional poverty measures (Foster, Greer and Thorbecke, 1984) to the case of multiple deprivations. Three multidimensional poverty measures are employed in this paper.³ The multidimensional headcount ratio or H represents the proportion of households identified as multidimensionally poor, according to the parameter k or

$$H = \frac{q(k)}{n}, \tag{1}$$

where q(k) is the number of multidimensionally poor people (or households) and n is the total number of people. The intensity of multidimensional poverty or A is defined as the average share of weighted indicators in which poor households are deprived, written as

$$A = \sum_{i=1}^q \frac{c_i(k)}{q(k)}. \tag{2}$$

This measure captures the breadth of deprivations experienced by multidimensionally poor households. The score is *censored* because it ignores the deprivations experienced by households that are not multidimensionally poor. The adjusted headcount measure or M₀ is calculated as the product of H and A:

$$M_0 = HA = \frac{q}{n} \frac{1}{q} \sum_{i=1}^q c_i(k) = \frac{1}{n} \sum_{i=1}^q c_i(k). \tag{3}$$

M₀ measures the sum of the weighted deprivations experienced by poor households, divided by the maximum number of deprivations that could be experienced by all households if all households were poor and deprived in all dimensions. M₀ increases either when an additional household becomes multidimensionally poor or any poor household becomes deprived in another dimension.⁴

4. Methodology

4.1 Evaluation design and data

The evaluation employs a quasi-experimental pipeline design (White and Barbu, 2006) based on the ex post comparison of treatment group households with a control group composed of households that had been selected for the programme but had not yet received the treatment. This evaluation design was chosen

³ The other measures within the family of Alkire-Foster measures, M1 and M2, cannot be computed because all the indicators included in our measure are dichotomous.

⁴ M₀ satisfies both the poverty focus and the deprivation focus axioms: the measure of poverty does not change when there is a change in any achievement of a non-poor person, and it is sensitive to the amount of deprivations of the poor (Alkire et al., 2015).

because TECHO-Argentina did not want to alter in any way its housing programme's processes or delay the provision of a dwelling to households that satisfied the eligibility requirements and for whom resources were available. These conditions ruled out the possibility of using either an experimental design or the difference-in-difference methodology.

The treatment group contains the participant households that received the TECHO house during the first half of 2014, and the control group contains the households that received the house in the first half of 2015. The same participant selection process, described in section 2, was applied in each year. In both years all of the selected households chose to participate in the programme and signed a contract in which they agreed to pay approximately 10% of the cost of the materials, attend preconstruction meetings and prepare their plot of land prior to construction.⁵ Both the treatment and control group baseline surveys were conducted between one and four weeks before construction of the TECHO house and at the time of the baseline survey the household members already knew that they had been selected to participate in the programme.⁶ The treatment group follow-up survey was carried out during the first half of 2015, approximately one year after each household had received the TECHO house.⁷

The cases were selected from the 34 informal settlements in the Province of Buenos Aires, Argentina, in which TECHO carried out its emergency housing programme during 2014–15. The sample includes settlements in the northern, western and southern zones of Greater Buenos Aires and in Greater La Plata but does not include informal settlements in the City of Buenos Aires. A total of 17 neighbourhoods provided both treatment and control cases, 5 provided only treatment cases and 12 provided only control cases.⁸ All of the settlements satisfy TECHO's operational definition of informal settlement: neighbourhoods with at least eight families in which at least half of the households do not have a formal land title nor formal access to at least two basic services (water, sanitation and electricity) (TECHO, 2013). Although the majority of the neighbourhoods included in the study have access to the electrical network, most families have an illegal connection. Fifteen of the 34 neighbourhoods are connected to the public

⁵ A small number of households did not receive the dwelling because they did not comply with contract's conditions.

⁶ The sample size of the treatment group baseline survey was 334 and the size of the follow-up survey was 293. Of the 41 households lost to attrition, 18 sold the TECHO house and left the neighbourhood, 20 moved from the neighbourhood taking the TECHO house with them and 1 house burned down. Robustness tests presented in Mitchell, Macció and Mariño Fages (2016) indicate that attrition likely had a small effect on the evaluation results.

⁷ In a previous paper (Mitchell, Macció and Mariño Fages, 2016), we present estimates of the before-after difference for the treatment group. Although these estimates do not measure the programme's causal effects, they serve as descriptive measures of what happened to the treatment group.

⁸ Mitchell, Macció and Mariño Fages (2016) show that when the impact estimates were re-estimated based only on the cases from the 17 neighbourhoods from which both treatment and control group households were drawn, the results were virtually unchanged in terms of both statistical and economic significance.

water system (with irregular connections in all but three cases) and in the rest of the neighbourhoods the households depend on well water. The most common form of sewage disposal is cesspit.

Table 1: Pre-treatment characteristics of the treatment and control groups

Characteristics	Control		Treatment		t-stat	Normalized difference
	Mean	Std. Dev.	Mean	Std. Dev.		
Number of nuclear family units	1.38	0.60	1.37	0.61	0.25	-0.02
Respondent and spouse in household	0.75	0.44	0.70	0.46	1.34	-0.11
Share of members under age 5	0.27	0.19	0.26	0.19	0.14	-0.01
Share of members ages 6 to 17	0.24	0.21	0.26	0.21	-1.11	0.09
Pregnant woman in household	0.14	0.34	0.15	0.35	-0.36	0.03
Respondent's age	32.99	12.47	34.26	13.22	-1.25	0.10
Respondent is male	0.17	0.38	0.16	0.37	0.22	-0.02
Respondent is a foreigner	0.04	0.20	0.06	0.23	-0.88	0.07
Respondent is from another province	0.19	0.40	0.19	0.39	0.07	-0.01
Respondent completed only primary school	0.22	0.41	0.26	0.44	-1.10	0.09
Respondent has some secondary school	0.37	0.48	0.36	0.48	0.12	-0.01
Respondent completed secondary school	0.13	0.34	0.11	0.32	0.70	-0.06
Respondent works	0.44	0.50	0.47	0.50	-0.62	0.05
Respondent's spouse works	0.63	0.48	0.59	0.49	1.06	-0.08
Number of other family members who work	0.50	0.89	0.49	0.84	0.22	-0.02
Income per capita (4th-Q 2014 pesos)	1,000	700	1,019	873	-0.31	0.02
At least one member has asthma	0.22	0.42	0.22	0.42	-0.01	0.00
n	304		246			

Note: Data from TECHO Housing survey, 2014-15.

The household survey collected data on sociodemographic characteristics; income; the size, quality and use of the house; physical and psychological health; sleep; security; social relations and perception of wellbeing. The survey respondent was the person who spent the most time in the home and was usually the mother of the principal family unit. The survey also collected, through the use of a sketch, information on the location of the house on the land parcel, the location and use of each room, and the location of each bed and which family members sleep in them.

As both the treatment and the control groups were selected for participation in the programme using the same processes, one would expect them to be highly similar. Table 1 compares 17 pre-treatment characteristics of the treatment and control groups. The tests of differences in means indicate that there is not a statistically significant difference between groups for any of the variables considered. Furthermore, the value of the normalized difference tends to be quite low, providing further evidence of the similarity between the treatment and control groups (see Imbens, 2015).

We use the following regression equation to estimate the average effect of treatment on the treated for outcomes of interest, Y_i :^{9 10}

$$Y_i = \alpha + \beta X_i + \delta T_i + \varepsilon_i, \quad (4)$$

where T_i is an indicator of treatment, X_i is a set of observed household characteristics and ε_i is an error term reflecting unobserved characteristics that also affect Y_i . The set of observed household characteristics correspond to the variables presented in Table 1 as well as to neighbourhood dummy variables. All of the control variables are based on pre-treatment data for both the control and treatment groups. Robust standard errors with neighbourhood clustering are used for all of the regression models.

4.2 Multidimensional impact evaluation strategy

We use the Alkire-Foster method to construct a multidimensional poverty index with dimensions related to the built environment. The dimensions, indicators and weights chosen for our measure are presented in Table 2. The unit of analysis is the household. For the construction of the Alkire-Foster measure we chose to maintain the same dimensions of analysis as were used in previous work (Mitchell, Macció and Mariño Fages, 2016).¹¹ The chosen dimensions are physical health, psychological health, sleep, privacy, interpersonal relations and security. These dimensions were selected based on TECHO's experience working with households living in informal settlements and a review of the literature. An effort was made to include indicators that correspond to individual (or household) functionings—what people can be or do—rather than income or other means for achieving wellbeing.

The first two dimensions of analysis—physical and psychological health—are those for which there has been the greatest amount of research with respect to their relationship to housing. The health literature has tended to focus on the pathological effects of some aspect of housing, such as poor ventilation,

⁹ See Angrist and Pischke (2009) for an explanation of why an ordinary least squares regression model is appropriate for measuring average treatment effects on binary outcome variables.

¹⁰ Mitchell, Macció and Mariño Fages (2016) employ two different strategies for measuring impact (a before-after comparison and an ex post treatment-control comparison). The study also uses both regression analysis and the propensity score matching technique and presents the results of the following robustness tests: non-random attrition, neighbourhood differences, the division of the household or change in the respondent between the baseline and follow-up surveys, seasonal differences in the time of the survey, alternative matching procedures, use of contemporaneous regression controls instead of pretreatment values and a falsification test in which treatment group post-treatment outcomes were replaced with the values from the baseline survey.

¹¹ The only dimension that was included in Mitchell, Macció and Mariño Fages (2016) but left out here is 'activities in the dwelling', which had two indicators: 'dwelling has a tranquil place for children to study' and 'at least one member practices a trade or handicraft in the dwelling'. We decided to exclude the first indicator from the Alkire-Foster measure because there were no survey responses for households without school-age children and the second indicator because factor analysis showed an inverse relationship between that variable and the rest of the deprivation measures.

Table 2: Multidimensional Poverty Index Dimensions, Indicators and Weights

Dimension	Indicators	Definition of the indicator	Weight (%)
Physical health			16.67
	Respiratory health	At least one member had frequent cough or congestion during the past year	8.33
	Joint pain	At least one member had joint pain during the past month	8.33
Psychological health			16.67
	Emotional health	Respondent had negative feelings (anxiety, depression, etc.) during past month	8.33
	Stress due to conflict	Respondent felt stress due to conflict in the home during the past month	8.33
Sleep			16.67
	Sleep quality	At least one member wakes up frequently at night	8.33
	Sleep quantity	Respondent sleeps less than 6 hours per night	8.33
Privacy			16.67
	Privacy within the home	Household members do not have space to dress in private	8.33
	Sleep privacy	At least one bed is overcrowded	8.33
Interpersonal relations			16.67
	Social relations within the household	Frequent conflict in the home due to lack of space	8.33
	Social relations with people from outside the household	Household never receives friends or relatives in home or the respondent feels uncomfortable doing so	8.33
Security			16.67
	Physical safety	Respondent fears that the dwelling could collapse.	8.33
	Security from theft	Household's dwelling suffered a robbery during the last year	8.33

insecure construction or the use of inadequate or toxic materials (Newman, 2008; Thomson et al., 2001). Research has demonstrated, for example, the association between humidity, mould and deficient heating on respiratory problems in children (Peat and Dickerson, 1998; Brunekreef, et al, 1989; Benicio et al., 2004). Overcrowding, moreover, fosters the spread of disease (Fonseca et al 1996; Murtagh et al 1993). Recent research also has provided evidence on the causal effect of housing programmes on health. Cattaneo et al. (2007), for example, showed that a Mexican programme that replaced dirt with cement floors reduced the incidence of parasites, diarrhoea and anaemia, and Galiani et al. (2016) demonstrated the effect of emergency housing programmes in El Salvador and Mexico on reducing diarrhoea in children. Housing can affect mental health both directly, due to poor ventilation, noise and exposure to light, and

indirectly, by reducing people's perceptions of the degree of control they have over their lives (Evans, 2003; Krieger and Higgins, 2002).

The third dimension, sleep, has been shown to be associated with both housing size and the quality of construction (Cardinali et al., 2014; Simonelli et al., 2015). Housing that does not provide adequate insulation from noise, heat, cold and wind can interrupt sleep or even force people to get up at night. Simonelli et al. (2013) conclude that emergency housing improves the subjective sleep quality and decreases sleep disturbances and daytime somnolence. Moreover, chronic health problems such as cardiovascular disease, diabetes, depression and obesity occur more frequently among persons who do not get adequate sleep (Buxton & Marcelli, 2010; Strine and Chapman, 2005).

The fourth dimension, privacy, is closely related to dwelling size. Overcrowding, it is argued, increases social demands, provokes resource competition and negatively affects psychological health when people feel that their daily activities are observed by others (Gove, Hughes and Galle, 1979). Overcrowded housing is also associated with bed overcrowding (Cardinali et al., 2014).

The fifth dimension is interpersonal relations. Both overcrowding and the stress and stigma of living in poor quality housing can generate conflict between household members and impede the development of supportive interpersonal relations both within the household and with people from outside of the home (Evans, 2003; Gove, Hughes and Galle, 1979).

The sixth dimension is security, defined as the absence of risk or threat. Inadequate building design can produce risks to physical safety, allow the intrusion of external threats and diminish perceptions of security and control (Bonnefoy, 2007). Studies carried out in Brazil and South Africa have demonstrated that injuries associated with precarious housing are higher in informal settlements than in other urban neighbourhoods (Bartlett, 2002). Housing improvements can reduce both the risk of injury and the incidence of robbery by making the house more impervious to theft.

In order to construct a parsimonious measure that would be easier to analyse and interpret, we chose to select only two indicators within each dimension from our original list of 31 indicators. To avoid selecting only those indicators for which we had found a statistically significant effect (a problem of 'cherry picking'), we established the following theoretical-methodological and empirical criteria to guide the selection.

- 1) Cover the diverse types of deprivations within each dimension (for example, include indicators of both sleep quantity and quality).
- 2) Include the indicators most critical to household wellbeing (for example, in the privacy dimension, include a measure of 'lack of space to change in private', but exclude 'lack of space to store belongings').

- 3) Select the indicators that appear to be measured more accurately (for example, include ‘lack of space to change in private’ instead of ‘lack of space to be alone’).
- 4) Use the relatively less demanding threshold when more than one had previously been considered (for example, include ‘respondent had negative feelings during the previous month’ instead of ‘respondent had negative feelings almost every day during the previous month’).
- 5) Prioritize objective over subjective indicators.¹²
- 6) Not include variables with a large number of missing values.¹³
- 7) Exclude variables defined for a subgroup of households.

Following these criteria, we reduced the number of impact variables to 12 indicators in six dimensions. In some cases, the indicator refers to the collective deprivation of the household (for example, ‘frequent conflict in the home due to lack of space’) and in other cases the indicator aggregates the situations of deprivation of the household members (for example, ‘at least one member frequently wakes up at night’). The measures of emotional health and sleep quantity refer to the situation of the respondent, the only household member for whom these data were collected. Equal weights are given to all dimensions and to all indicators within each dimension. This decision was made on the basis that all of the dimensions are considered to be essential aspects of housing-related wellbeing.

In order to evaluate the associations between our 12 housing-related indicators and identify possible redundancy, we produced 66 contingency tables to assess the associations between all pairs of indicators. For each table we obtained the χ^2 statistic for testing the independence between dichotomous variables and calculated the redundancy coefficient, R.¹⁴ We used the baseline data for the treatment and control groups and the censored headcounts. Following Alkire et al. (2015), we then analysed the redundancy coefficient for those pairs of indicators that are not independent according to the χ^2 test ($p < 0.05$). The null hypothesis of independence was rejected for only 15 of the 66 tests and only 4 of these 15 pairs have values of R greater than 0.75. These pairs are (i) sleep privacy and privacy within the home; (ii) sleep privacy and emotional health; (iii) emotional health and sleep quantity and (iv) emotional health and conflicts due to lack of space.

¹² The preference for objective over subjective measures is motivated by what Sen (1985) calls ‘adaptive preferences’, whereby people who experience prolonged periods of deprivation become satisfied with low levels of objective wellbeing. An additional motivation is to reduce the potential problem of ‘confirmation bias’ that could arise because TECHO volunteers carried out the survey field work.

¹³ The Alkire-Foster measure requires that each observation has a valid value for all of the indicators, potentially causing a large loss of observations even if each indicator is missing information for a very small number of households.

¹⁴ The coefficient R is equal to the proportion of observations of deprivation in both variables, which represents the joint deprivation, relative to the minimum of the two censored or uncensored headcount ratios (Alkire et al., 2015).

Based on both empirical and theoretical considerations we decided to maintain the selected indicators in our Alkire-Foster measure. In empirical terms, none of the indicators has a high level of association with all of the remaining indicators. In theoretical terms, all of the chosen indicators measure aspects of deprivation considered to be important for understanding how the TECHO programme affects multidimensional poverty.

Table 3 presents the deprivation rates for the 12 indicators included in the Alkire-Foster measure based on the treatment and control group baseline data. The t-test of difference in means indicates that there is a statistically significant difference between groups only in the case of two indicators: joint pain and social relations within the household. In each of these indicators the incidence of deprivation is larger in the control than in the treatment group.

We set the poverty threshold, k , at 33% or four weighted indicators. This is the same poverty threshold used in the global Multidimensional Poverty Index (Alkire and Santos, 2011). For a household to be classified as multidimensionally poor it would need to be deprived in two full dimensions or be partly deprived in four dimensions. Section 5.3 presents robustness tests on the poverty threshold.

Multidimensional poverty in dimensions related to the built environment can be interpreted as an accumulation of housing-related deprivations that surpasses a critical limit. It is reasonable to imagine that when a household faces a limited number of deprivations, a functioning in one dimension could help the household members to cope with deprivations in other dimensions. For example, strong interpersonal relations and sound psychological health could enable family members to deal with the lack of privacy, frequent sleep interruptions and physical insecurity associated with living in an inadequate dwelling. In this case deprivations in three indicators would not constitute a situation of multidimensional poverty. However, if a family member were to become ill or the household were to suffer a robbery, it is reasonable to imagine that these situations could cause the household to reach a critical state of destitution. A household in this state would be considered to be multidimensionally poor.

The Alkire-Foster multidimensional poverty method enables us to obtain an estimate of the programme's effect on the joint distribution of deprivations. We employ the following three regression equations, which are based on equation (4), to measure the effect of TECHO's programme on the three Alkire-Foster measures. Equation (5) is used to measure the effect on the incidence of multidimensional poverty, where h_i is a dichotomous variable =1 if household i is multidimensionally poor and = 0 otherwise.

$$H: h_i = \alpha + \beta X_i + \delta_1 T_i + \varepsilon_i \quad (5)$$

Table 3: Comparison of control and treatment group baseline deprivation rates

Dimension/Indicator	Control	Baseline Treatment	Difference	n
<i>PHYSICAL HEALTH</i>				
Respiratory health	0.368 (0.483)	0.318 (0.467)	-0.050	549
Joint pain	0.319 (0.467)	0.252 (0.435)	-0.067	* 550
<i>PSYCHOLOGICAL HEALTH</i>				
Emotional health	0.668 (0.472)	0.634 (0.483)	-0.034	547
Stress due to conflict	0.438 (0.497)	0.459 (0.499)	0.021	546
<i>SLEEP</i>				
Sleep quality	0.461 (0.499)	0.426 (0.496)	-0.034	548
Sleep quantity	0.250 (0.434)	0.216 (0.412)	-0.034	545
<i>PRIVACY</i>				
Privacy within the home	0.405 (0.492)	0.382 (0.487)	-0.023	545
Sleep privacy	0.714 (0.453)	0.686 (0.465)	-0.028	549
<i>INTERPERSONAL RELATIONS</i>				
Social relations within	0.289 (0.454)	0.203 (0.403)	-0.087	** 541
Social relations outside	0.359 (0.48)	0.354 (0.479)	-0.005	547
<i>SECURITY</i>				
Physical safety	0.385 (0.487)	0.363 (0.482)	-0.022	549
Security from theft	0.204 (0.404)	0.167 (0.373)	-0.037	550

Notes: Data from TECHO Housing survey, 2014-15. Standard errors in parenthesis.

**p < 0.05;

*p < 0.1.

Equation (6) is used to measure the programme’s effect on the intensity of poverty among the multidimensionally poor. The dependent variable is the deprivation score and note that the regression is estimated for the subsample of poor households.

$$A: c_i(k) = \alpha + \beta X_i + \delta_2 T_i + \varepsilon_i, \quad c_i \geq k \tag{6}$$

Equation (7) is used to measure the programme’s effect on the adjusted headcount measure M_0 . The deprivation score is the dependent variable again in this model, but in this case the regression is estimated for the full sample (poor and non-poor households).

$$M_0: c_i(k) = \alpha + \beta X_i + \delta_3 T_i + \varepsilon_i \tag{7}$$

The coefficient δ_1 gives an estimate of the effect of treatment on H, δ_2 an estimate of the effect of treatment on A and δ_3 an estimate of the effect of treatment on M_0 . Robust standard errors with neighbourhood clustering are used for all of the regression models.

5. Evaluation results

The analysis of the effects of TECHO's emergency housing programme on multidimensional poverty is divided into five subsections. The first presents the estimated effects on individual outcome measures and the second the effects on the Alkire-Foster multidimensional poverty measures. The third subsection analyses the robustness of the results to changes in the process of selecting of indicators, the indicator weights and the poverty threshold. The fourth subsection assesses the heterogeneity of treatment effects on initially poor versus non-poor households. The last subsection analyses the contribution of the indicators and dimensions to overall poverty.

5.1 Effects on individual outcomes

Before analysing the effects of treatment on the impact indicators it is important to consider the direct effects of the programme on the size and quality of the dwelling (see Table A1 in the Annex). TECHO's programme caused the average size of the dwelling to increase from 29.5 to 36.5 m² and the average number of rooms to increase from 2 to 2.7.¹⁵ As a result, the percentage of households with less than 5 m² per person dropped from 47% to 23% and the percentage with more than three persons per room from 37% to 13%.¹⁶ With regard to housing quality, the programme did not have a statistically significant effect on the share of rooms with a roof or walls made of materials classified as good quality. This result is likely due to the fact that most of the original dwellings were made of components that in theory were appropriate but were of such poor quality that they did not provide adequate protection (such as, walls made of discarded wood or uninsulated tin roofs). The programme did, however, produce a large and statistically significant reduction in the average share of rooms that had *problems* with the roof and walls.¹⁷ In addition, the share of rooms that flood when it rains fell from 47% to 22% and the share that are very or quite humid fell from 66% to 28%. Galiani et al. (2016) also found that TECHO's housing programme produced a statistically significant improvement in housing quality and the magnitude of the effect was

15 The effect of the programme on the size and quality of housing varies according to the use the household gives to the dwelling. Around a third of treatment group households replaced their dwelling with the TECHO house and two-thirds used it as an additional room or as a dwelling for one the household's nuclear family units.

16 Although a large proportion of households experienced some change in the household composition (such as births, deaths, changes in spouses, etc.), the programme did not produce a statistically significant change in the average household size.

17 While the change in the share of rooms with dirt floors or with dirt or water seepages in the floor also declined according to the Holm procedure, this effect was not statistically significant at the 10% level.

larger in El Salvador, the country with the worst initial housing conditions. In that study, however, after adjusting for multiple hypothesis testing, the estimated effect on the number of rooms was statistically significant only in the case of Mexico. Finally, consistent with previous research (Galiani and Schargrodsky, 2015; Galiani et al, 2016), the programme did not produce a statistically significant effect on household size.

Table 4 presents the estimates of the treatment's effect on each of the variables included in our Alkire-Foster measure based on the regression equation specified in equation (4).¹⁸ Column 4 presents the coefficient estimates for treatment for the full sample and columns 5 through 7 the results of the analysis of the heterogeneity of the treatment effects (see section 5.4). Figure 1 compares the values of the outcome indicators for the control group and the treatment group based on the estimated treatment effects. Tables 4 also presents the results of the Holm stepwise procedure used to correct for the increased likelihood of type I error (rejection of a true null hypothesis) when carrying out multiple hypothesis tests (Lehmann and Romano, 2005).¹⁹

Physical health and sleep are the dimensions for which the estimated treatment effects are weakest. While the percentage of households with highly frequent cough or congestion declines from 37% to 26%, based on the Holm procedure for multiple hypothesis tests, this effect is not statistically significant at the 10% level. There is also no statistically significant effect on the percentage of households in which at least one member had joint pain during the past month.²⁰ In the sleep dimension, the results indicate that the TECHO house reduced the percentage of households in which at least one member wakes up frequently at night from 46% to 29%, but based on the Holm procedure the effect is not statistically significant. There also is no effect on the percentage of respondents that sleep less than six hours per night.²¹

In the psychological health dimension, the effects of the emergency housing programme are mixed. The programme produces a large reduction in the percentage of households in which the respondent had negative feelings (like sadness, anxiety or depression) during the past week, which falls from a control

18 The small differences in the treatment effect coefficients with respect to the results presented in Mitchell, Macció and Mariño Fages (2016) are due to differences in sample size, as the estimates presented here are based on the cases that have valid values for all of the indicators included in the Alkire-Foster measure. All of the substantive results are unchanged.

19 To apply the Holm procedure, one first orders the estimation results by p-values i from smallest to largest and then sequentially rejects each hypothesis as long as the p-value $< \alpha_i = \alpha / (s - i + 1)$. We employ the Holm stepwise adjustment instead of the Bonferroni procedure because the latter is considered to increase the likelihood of type II errors too much (failure to reject a false null hypothesis) (Perneger, 1998).

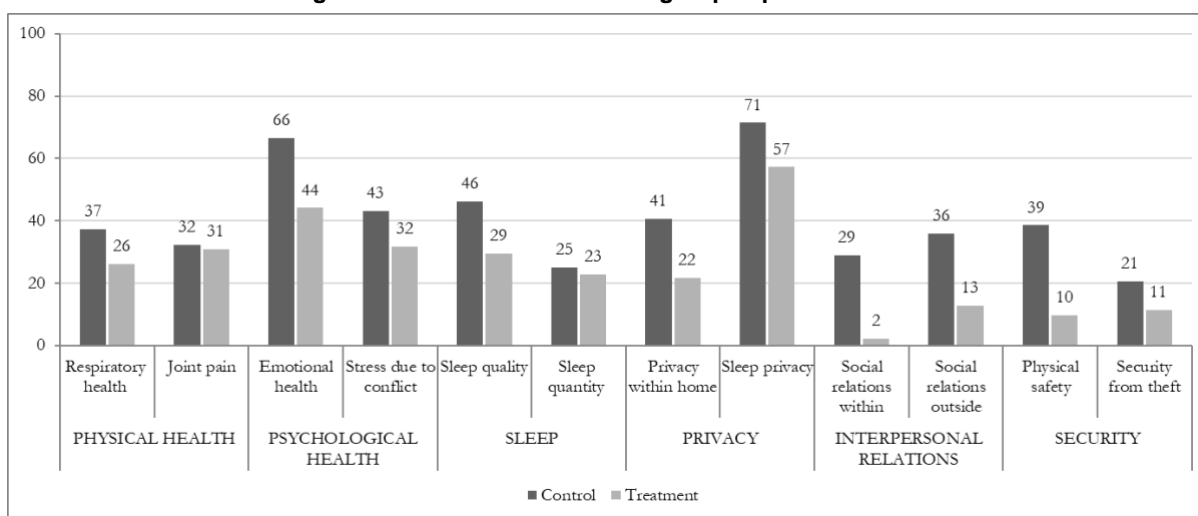
20 Galiani et al. (2016) also did not find a statistically significant effect on the incidence of respiratory health in children.

21 Simonelli et al. (2013), in contrast, found that the programme had a positive effect on sleep quality. The difference in the results may be because that study was based on a before-after comparison and the outcome variable was a composite index of sleep quality composed of 19 self-rated questions.

group mean of 66% to 44%. The percentage of respondents that feel stress due to conflict in the home also declines, but the effect is not statistically significant after making the Holm procedure adjustment.

In both the privacy and interpersonal relations dimensions, we find that the programme has a large and statistically significant effect on all indicators. The percentage of households in which the members do not have space to dress in private declines by almost half from 41% to 22%, and the percentage of households in which at least one bed is overcrowded declines from 71% to 57%. In addition, the percentage of households that have frequent conflict due to lack of space falls from 29% to just 2%, and the percentage of households that either never receive friends or neighbours in the home or feel quite or very uncomfortable doing so falls from 36% to 13%.

Figure 1: Treatment and control group deprivation rates



Notes: Data from TECHO Housing survey, 2014-15.

The TECHO programme also reduced both the perception of insecurity and the incidence of robbery. The percentage of respondents who are concerned that the building could collapse declines from 39% to 10%, and the percentage of households that experienced a robbery during the past year falls by half from 21% to 11%. It is relevant to note that the impact evaluations carried out in El Salvador, Mexico and Uruguay showed that the TECHO programme had a positive effect on the perception of security but did not reduce the incidence of robbery (Galiani et al., 2016).

Table 4: Treatment effect on the Alkire-Foster measure indicators

Dimension/Indicator	Control Mean (Std. Dev.)	Estimated treatment effect (Full sample)	Regression results with heterogeneous effects of treatment for poor and non-poor households						
			T	H	T*H	T	H	T*H	
<i>PHYSICAL HEALTH</i>									
Respiratory health	0.372 (0.484)	-0.111 (0.053)	**	†	0.045 (0.063)	0.346 (0.044)	***	-0.218 (0.071)	***
Joint pain	0.322 (0.468)	-0.014 (0.059)			0.049 (0.076)	0.315 (0.044)	***	-0.105 (0.074)	
<i>PSYCHOLOGICAL</i>									
Emotional health	0.664 (0.473)	-0.224 (0.063)	***		-0.036 (0.09)	0.489 (0.057)	***	-0.243 (0.09)	**
Stress due to conflict	0.432 (0.496)	-0.116 (0.06)	*	†	0.166 (0.085)	* † 0.483 (0.044)	***	-0.395 (0.08)	***
<i>SLEEP</i>									
Sleep quality	0.462 (0.499)	-0.168 (0.08)	**	†	0.001 (0.087)	0.370 (0.067)	***	-0.253 (0.088)	***
Sleep quantity	0.249 (0.433)	-0.022 (0.053)			0.110 (0.069)	0.199 (0.049)	***	-0.194 (0.057)	***
<i>PRIVACY</i>									
Privacy within home	0.405 (0.492)	-0.188 (0.056)	***		0.024 (0.106)	0.301 (0.067)	***	-0.304 (0.106)	***
Sleep privacy	0.714 (0.453)	-0.142 (0.042)	***		-0.084 (0.06)	0.129 (0.043)	***	-0.072 (0.071)	
<i>INTERPERSONAL RELATIONS</i>									
Social relations within	0.289 (0.454)	-0.267 (0.057)	***		-0.140 (0.072)	* † 0.259 (0.04)	***	-0.198 (0.051)	***
Social relations outside	0.359 (0.48)	-0.232 (0.03)	***		-0.086 (0.042)	** † 0.252 (0.058)	***	-0.198 (0.062)	***
<i>SECURITY</i>									
Physical safety	0.385 (0.487)	-0.288 (0.043)	***		-0.096 (0.055)	* † 0.293 (0.064)	***	-0.279 (0.069)	***
Security from theft	0.206 (0.405)	-0.092 (0.035)	**		-0.024 (0.059)	0.111 (0.045)	**	-0.099 (0.073)	
n	301	546				525			

Notes: Data from TECHO Housing survey, 2014-15. Standard errors in parenthesis, except where indicated.

***p < 0.01;

**p < 0.05;

*p < 0.1.

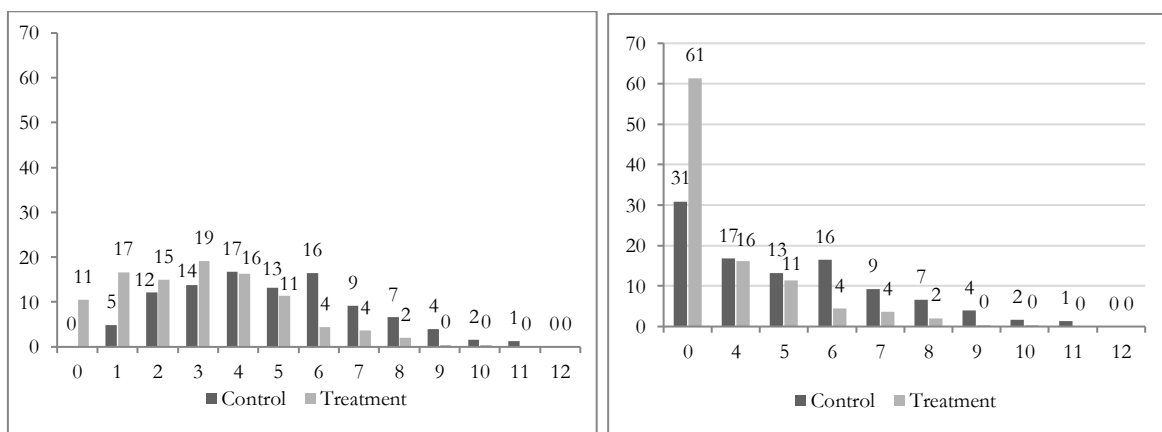
† indicates that the variable is not statistically significant at the 10% level according to the Holm procedure for multiple hypothesis testing.

5.2 Effects on multidimensional poverty

The analysis of the treatment effects on individual outcome measures showed that the TECHO programme reduced deprivation in multiple wellbeing dimensions. We begin the analysis of the programme’s effect on multidimensional poverty by comparing the distributions of the deprivation score, c_i , and the censored deprivation score, $c_i(k)$, for treatment and control households (Figure 2). The graph on the left clearly illustrates the leftward shift in the distribution of the number of deprivations experienced by treatment group households compared with control group households. The censoring of the deprivation score at the poverty threshold k shows that the fraction of treatment group households with less than four deprivations doubles the fraction corresponding to control group households. These comparisons provide initial evidence of the program’s effect on both the incidence of multidimensional poverty and on the intensity of poverty experienced by the multidimensionally poor.

Table 5 presents the results of the estimation of the effects of the TECHO programme on the three Alkire-Foster measures of multidimensional poverty. The table first presents the control group baseline estimate of each measure, followed by the estimated treatment effect based on the estimation of regression equations (5) through (7), followed by the estimated value of each measure for treated households (calculated as the sum of the control group mean and the treatment effect) and the observed value of each measure for treatment group households.

Figure 2: Distributions of c_i and $c_i(k)$ of treatment and control group households



Notes: Data from TECHO Housing survey, 2014-15.

The results indicate that the TECHO programme caused a statistically significant reduction in all three of the Alkire-Foster measures at the 1% significance level after adjusting for multiple hypothesis testing using the Holm procedure. The multidimensional headcount fell by approximately 30 percentage points. Whereas 69% of control group households were multidimensionally deprived in dimensions related to housing, the percentage fell to 39% one year after the treatment group households had received the TECHO house.

Table 5: Treatment effect on the Alkire-Foster measure for k=33%

	H	A	M ₀
Control group	0.691	0.501	0.346
Treatment effect (Standard error)	-0.306 *** (0.05)	-0.080 *** (0.016)	-0.183 *** (0.024)
Treated group (estimated)	0.385	0.421	0.163
Treated group (observed)	0.384	0.427	0.164
Diff. Treated(Obs.)-Control	-0.307	-0.074	-0.182
Shapley decomposition			
Incidence of poverty effect	-0.142		
Intensity of poverty effect		-0.040	
Change in M ₀			-0.182
n	546	302	546

Notes: Data from TECHO Housing survey, 2014-15.

***p < 0.01.

Poverty intensity or A declined by eight percentage points. Whereas at baseline the multidimensionally poor were, on average, deprived in half of the weighted indicators, after treatment this proportion fell to 42%.²² The decline in both the incidence and the intensity of poverty caused the adjusted headcount ratio to drop by more than half from .346 to .163. This means that multidimensionally poor households went from being deprived in a third of all of the indicators in which all households can be deprived to only experiencing 16.3% of these deprivations. Following Roche (2013), Table 5 also presents the Shapley decomposition of the effect of treatment on M₀ into two components: the incidence of poverty and the intensity of poverty among the multidimensionally poor.

While this procedure is usually applied to the analysis of variations over time, it can also be used for cross-sectional analysis (Alkire et al., 2015). The incidence of poverty effect is equal to the average poverty intensity of the two groups multiplied by the difference in the headcount ratio between groups. The intensity of poverty effect is equal to the average of the headcount ratios of the two groups multiplied by the difference in the poverty intensity for the two groups.²³ The results indicate that for the chosen poverty

22 When interpreting the effect of treatment on A it is important to recall that, since poverty intensity is equal to the average deprivation score of poor households, when the multidimensional headcount also declines, the share of weighted deprivations is averaged over a smaller and poorer proportion of households. For this reason, A tends to change relatively slowly.

23 The decomposition equation adapted from equation 9.18 in Alkire et al. (2015) is

$$\Delta M_0 = \frac{A^T + A^C}{2}(H^T - H^C) + \frac{H^T + H^C}{2}(A^T - A^C).$$

The first term is the incidence effect and the second the intensity effect.

threshold of 33%, approximately three-fourths of the treatment effect on M_0 is explained by the decline in the incidence of poverty and one-fourth by the decline in poverty intensity.

5.3 Robustness analysis

In order to evaluate the robustness of the results to changes in the weighting and indicator selection criteria used to construct the Alkire-Foster measures, we carried out three different types of tests. In the first set of tests, we altered the weights used to construct the Alkire-Foster measures by increasing to 25% the weight assigned to one particular dimension and distributing equally the remaining 75% to the rest of the dimensions. This test was repeated for each of the six wellbeing dimensions. In the second set of tests, we altered the criteria used for the selection of indicators. For the first test we randomly selected two indicators within each dimension and for the second we randomly selected 12 indicators from the original list of 31 indicators.²⁴ For each test we re-estimated the effect of treatment on each of the three Alkire-Foster measures using regression models (5) – (7). The results are presented in Table 6.

The effects of the TECHO programme on multidimensional poverty are highly robust to variations in the weighting structure used to construct the Alkire-Foster measure. While the size of the treatment effect varies somewhat when alternative weighting structures are used,²⁵ the coefficient on treatment is statistically significant at the 1% level for all of the tests on all three poverty measures. The results are also quite robust to changes in the selection of indicators. For the case in which the indicators were randomly selected within each dimension, the treatment effect continues to be statistically significant for all three Alkire-Foster measures and the size of the effect on M_0 falls by only about 5%. When the 12 indicators are selected randomly over all dimensions, the effect of treatment on H and M_0 continues to be statistically significant at the 1% level. The magnitude of the effect on M_0 is about 60% of the initial value. With a random selection of indicators over all dimensions, the treatment effect on A is no longer statistically significant. Based on these robustness tests, we can conclude that the finding that TECHO's emergency housing programme produces a large and statistically significant reduction in multidimensionally poverty holds up to variations in the criteria used to construct the multidimensional poverty measures.

²⁴ It should be noted that, by chance, the 12 randomly selected indicators included at least one indicator from each dimension.

²⁵ For example, the effect on H varies from 28 percentage points, when a higher weight is given to the physical health dimension, to 34 percentage points, when a higher weight is given to the interpersonal relations dimension.

Table 6: Robustness tests on the criteria used to construct the Alkire-Foster measures

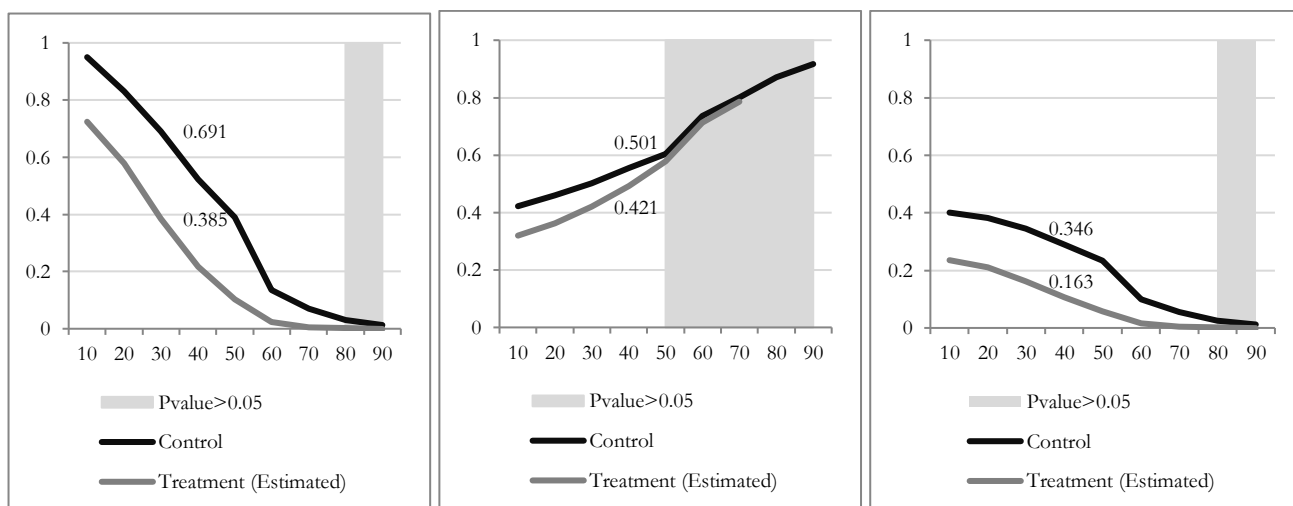
	Treatment effect	Standard error	t-stat	P-value	n
Dimension weights robustness tests					
Multidimensional headcount (H):					
Original Alkire-Foster measure	-0.306	0.050	-6.15	0.00	546
25% Privacy	-0.298	0.038	-7.84	0.00	546
25% Sleep	-0.296	0.046	-6.48	0.00	546
25% Physical health	-0.277	0.041	-6.71	0.00	546
25% Psychological health	-0.303	0.053	-5.67	0.00	546
25% Interpersonal relations	-0.341	0.039	-8.71	0.00	546
25% Security	-0.320	0.045	-7.10	0.00	546
Poverty intensity (A):					
Original Alkire-Foster measure	-0.080	0.016	-5.03	0.00	302
25% Privacy	-0.082	0.017	-4.85	0.00	279
25% Sleep	-0.062	0.017	-3.57	0.00	257
25% Physical health	-0.075	0.018	-4.14	0.00	253
25% Psychological health	-0.064	0.016	-4.07	0.00	281
25% Interpersonal relations	-0.074	0.016	-4.64	0.00	240
25% Security	-0.075	0.016	-4.83	0.00	239
Adjusted headcount (M ₀):					
Original Alkire-Foster measure	-0.183	0.024	-7.57	0.00	546
25% Privacy	-0.186	0.021	-8.71	0.00	546
25% Sleep	-0.175	0.024	-7.26	0.00	546
25% Physical health	-0.170	0.023	-7.45	0.00	546
25% Psychological health	-0.185	0.028	-6.61	0.00	546
25% Interpersonal relations	-0.197	0.021	-9.36	0.00	546
25% Security	-0.187	0.023	-8.17	0.00	546
Indicator selection robustness tests					
Multidimensional headcount (H):					
Original Alkire-Foster measure	-0.306	0.050	-6.15	0.00	546
Random selection within dimensions	-0.367	0.039	-9.34	0.00	562
Random selection over all dimensions	-0.238	0.046	-5.21	0.00	546
Poverty intensity (A):					
Original Alkire-Foster measure	-0.080	0.016	-5.03	0.00	302
Random selection within dimensions	-0.056	0.013	-4.26	0.00	183
Random selection over all dimensions	-0.019	0.012	-1.62	0.11	229
Adjusted headcount (M ₀):					
Original Alkire-Foster measure	-0.183	0.024	-7.57	0.00	546
Random selection within dimensions	-0.173	0.018	-9.39	0.00	562
Random selection over all dimensions	-0.107	0.020	-5.25	0.00	546

Note: Data from TECHO Housing survey, 2014-15.

In order to test the robustness of the results to changes in the poverty threshold, Figure 3 presents comparisons of the three Alkire-Foster measures corresponding to the control and treatment groups over a range of possible values of k. Note that the estimates of each measure for the treatment group are

calculated as the sum of the control group mean and the treatment effect using, respectively, regression equations 5 through 7. The shaded area of each graph indicates the values of k over which the effect of treatment on the Alkire-Foster measure is *not* statically significant at the 5% significance level. The results show that the effect of the TECHO programme on the multidimensional headcount ratio H is statistically significant for all values of the poverty cut-off less than or equal to 80%, indicating that the effect on the incidence of multidimensional poverty is not sensitive to the chosen poverty threshold. In the case of the poverty intensity, the graph shows that A is always higher for the control than the treatment group households and that the difference between groups is statistically significant only for values of k below 50%. Finally, M_0 is higher for the control than the treatment groups for all possible values of k , and the effect is statistically significant for all values of $k \leq 80\%$. Moreover, the difference is particularly large for values of $k \leq 50\%$. From this analysis it can be concluded that the program’s effect on reducing multidimensional poverty is highly robust to variations in the poverty threshold.

Figure 3: Effect of treatment on H , A and M_0 with variation in k



Notes: Data from TECHO Housing survey, 2014-15

5.4 Heterogeneity of treatment effects

Impact evaluations frequently assess the heterogeneity of the effects of treatment on programme participants with different initial characteristics. In this section we use the classification of the treatment and control group households as multidimensionally poor *before* receiving the TECHO house (at baseline) to analyse whether or not the programme has a relatively larger effect on those households that were initially multidimensionally poor.

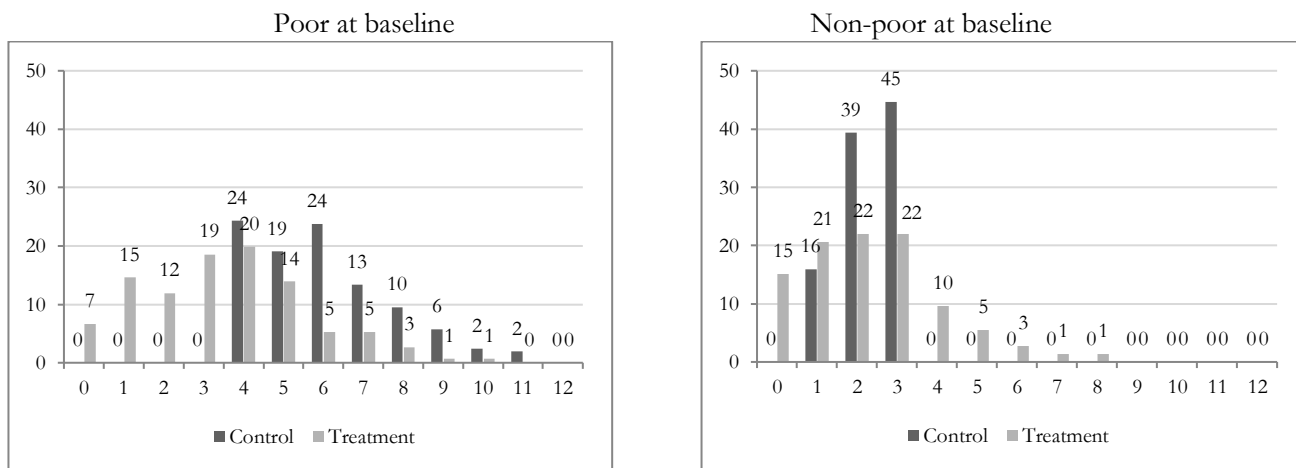
First, in order to analyse the heterogeneity of the treatment effects on the individual outcome measures, we added the following two additional explanatory variables to regression equation (4):

- (i) a dummy variable for multidimensional poverty, $h_i = 1$ if multidimensionally poor, otherwise $= 0$;
- (ii) and an interaction term between h and the indicator of treatment, T .

The results are presented in columns 5 through 7 of Table 4. The coefficient on T in the new regression equation provides an estimate of the effect of treatment on households that are initially not multidimensionally poor and the coefficient on the interaction term provides an estimate of the differential effect of treatment on households that were initially multidimensionally poor.

The results indicate that for all of the indicators in which we had initially found a statistically significant effect of treatment, the coefficient on the interaction term is negative and statistically significant. This means that the magnitude of the treatment effect is larger for households that were initially multidimensionally poor than for the full sample. Additional evidence of the heterogeneity of the treatment effect is that, after applying the Holm procedure for multiple hypothesis tests, we cannot reject the null hypothesis of no treatment effect on the initially non-poor households for any of the outcome variables considered.

Figure 4: Distribution of the deprivation score for control and treatment group households by initial poverty status



Note: Data from TECHO Housing survey, 2014-15.

Second, we use the distribution of the deprivation score of the control and treatment group households disaggregated according to *initial* poverty status (Figure 4) to assess the heterogeneity of the effects on the distribution of deprivations. On the one hand, when one considers only those households that were multidimensionally poor at baseline, it is evident that the treatment produces a leftward shift in the distribution of deprivation scores that is even larger than for the full sample (Figure 2). On the other hand, the distribution of the deprivation score for households classified as non-poor at the baseline shows that around 20% of the treatment group households became multidimensionally poor after receiving the TECHO house. One cannot conclude that this transition into multidimensional poverty was caused by the TECHO programme as many other factors could have contributed to the worsening of housing-related deprivation (for example, changes in the quality of access to public services or environmental conditions). It does, however, provide evidence of vulnerability to housing-related deprivations in informal

settlements. While *on average* the programme reduces both the incidence and the intensity of multidimensional poverty, some households experience a larger number of deprivations *after* receiving treatment.

5.5 Dimensional contributions to M_0

In order to analyse the extent of the contributions of each dimension to overall poverty, this section applies a dimensional breakdown to the adjusted headcount measure. M_0 can be expressed as the weighted sum of the censored deprivation rate

$$M_0 = \sum_{j=1}^d w_j h_j(k), \quad (8)$$

where d is the number of chosen dimensions, w_j is the relative weight assigned to the dimension and $h_j(k)$ is the proportion of households identified as multidimensionally poor and simultaneously deprived in dimension j (Alkire et al., 2015). The absolute contribution of the j th dimension to M_0 is $w_j h_j(k)$. Additionally, the percentage contribution of each dimension j to M_0 can be denoted by ϕ_j^0 , where

$$\phi_j^0(k) = w_j \frac{h_j}{M_0} \quad \text{for each } j = 1, \dots, d. \quad (9)$$

Table 7 first presents the control and ex-post treatment group deprivation rates for each of the indicators included in our Alkire-Foster measure and the absolute difference between the two groups.²⁶ The following columns show the control and treatment group censored deprivation rates for each indicator and the absolute difference between groups. Finally, equations (8) and (9) were used to calculate, respectively, the absolute and relative contributions of each indicator to M_0 . In each case we also calculate the absolute difference in the rates between the control and treatment groups and mark in grey the three values with the largest absolute differences. Also note that both the average of the censored deprivation rates and the sum of the absolute contributions are equal to M_0 .

The examination of the difference in the deprivation rates for the full sample indicates that physical safety, social relations between household members and social relations outside of the home are the indicators for which the programme produced the largest reductions in deprivation. When we focus our attention only the households that are simultaneously multidimensionally poor and deprived in each dimension (that is, we focus on the censored deprivation rates), we find that physical safety, sleep privacy and emotional health are the indicators for which there is the largest difference between the treatment and control groups. These same indicators have the largest absolute contribution to the change in multidimensional poverty

²⁶ Note that the absolute difference between these values is not equal to the treatment effects presented in Table 4, because the latter estimates were obtained by estimation of regression equation (4).

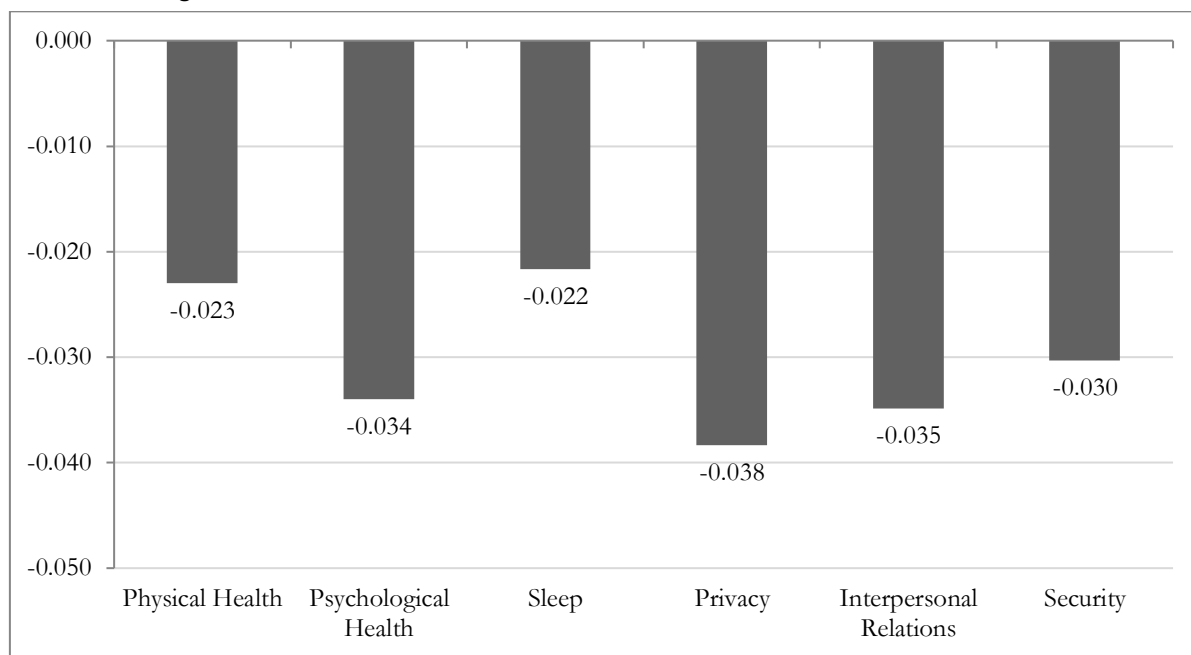
M0. The differences between the control and treatment group households in the censored deprivation score result from the combined effects of the movement out of multidimensional poverty by households that had previously been deprived in this indicator and the reduction in the overall rate of deprivation in the indicator.

Table 7: Percentage and Absolute Contributions to M0

	Deprivation rates (full sample)			Censored deprivation rates			Absolute contributions to M0			Percentage contributions to M0		
	Ctrl.	Treat.	Diff.	Ctrl.	Treat.	Diff.	Ctrl.	Treat.	Diff.	Ctrl.	Treat.	Diff.
<i>PHYSICAL HEALTH</i>												
Respiratory health	0.372	0.245	-0.127	0.336	0.176	-0.160	0.028	0.015	-0.013	0.081	0.089	0.008
Joint pain	0.322	0.278	-0.045	0.296	0.180	-0.116	0.025	0.015	-0.010	0.071	0.091	0.020
<i>PSYCHOLOGICAL HEALTH</i>												
Emotional health	0.665	0.498	-0.167	0.558	0.310	-0.248	0.047	0.026	-0.021	0.134	0.158	0.023
Stress due to conflict	0.432	0.318	-0.114	0.409	0.249	-0.160	0.034	0.021	-0.013	0.098	0.127	0.028
<i>SLEEP</i>												
Sleep quality	0.462	0.314	-0.148	0.402	0.229	-0.173	0.034	0.019	-0.014	0.097	0.116	0.019
Sleep quantity	0.249	0.184	-0.066	0.213	0.127	-0.086	0.018	0.011	-0.007	0.051	0.064	0.013
<i>PRIVACY</i>												
Privacy within home	0.405	0.245	-0.160	0.346	0.155	-0.190	0.029	0.013	-0.016	0.083	0.079	-0.004
Sleep privacy	0.714	0.576	-0.139	0.535	0.265	-0.270	0.045	0.022	-0.022	0.129	0.135	0.006
<i>INTERPERSONAL RELATIONS</i>												
Social relations within	0.289	0.061	-0.228	0.259	0.053	-0.206	0.022	0.004	-0.017	0.062	0.027	-0.035
Social relations outside	0.359	0.122	-0.236	0.302	0.090	-0.213	0.025	0.007	-0.018	0.073	0.046	-0.027
<i>SECURITY</i>												
Physical safety	0.385	0.057	-0.328	0.326	0.049	-0.277	0.027	0.004	-0.023	0.078	0.025	-0.054
Security from theft	0.206	0.114	-0.092	0.173	0.086	-0.087	0.014	0.007	-0.007	0.042	0.044	0.002
Average				0.346	0.164							
Sum							0.346	0.164	-0.182	1	1	0.000
n=546												

Note: Data from TECHO Housing survey, 2014-15.

Figure 6 presents the absolute difference in each dimension’s contribution to multidimensional poverty. Privacy was the dimension that contributed the most to the reduction in multidimensional poverty and within that dimension sleep privacy contributed relatively more than the measure of overall privacy within the home. Interpersonal relations was the next dimension that most contributed to the overall change in M0, followed by psychological health and then security. Within the security dimension, the contribution of the physical safety indicator is about three times larger than that of the indicator of security from theft. Finally sleep and health are the dimensions with the smallest contribution to the change in overall multidimensional poverty.

Figure 6: Absolute differences in each dimension's absolute contribution to M_0 

Note: Data from TECHO Housing survey, 2014-15.

6. Conclusions

This paper has sought to evaluate the effect of the NGO TECHO's emergency housing programme on the multidimensional deprivation experienced by households living in the informal settlements of Buenos Aires, Argentina. It employs an Alkire-Foster multidimensional poverty measure with dimensions related to the built environment and a quasi-experimental 'pipeline' evaluation design.

The paper demonstrates that the TECHO housing programme produces a large reduction in multidimensional poverty. The percentage of households identified as multidimensionally poor declines substantially, from 69% to 39%. The average share of indicators in which the multidimensionally poor are deprived declines from 50% to 42%. Together these effects cause the adjusted headcount measure to fall by more than half. Multidimensionally poor households go from experiencing deprivation in a third of all of the indicators in which all households can be deprived to only experiencing 16.5% of these deprivations. This result provides strong evidence that TECHO's programme contributes to what Wolff and de-Shalit (2007) call a 'de-clustering' of disadvantage. Furthermore, it is shown that these results are robust to variations in the selection of deprivation indicators, indicator weights and the poverty threshold.

The Alkire-Foster method also provided a framework for analysing the heterogeneity of the treatment effects on households with different levels of initial deprivation. The results show that the magnitude of the reductions in deprivations is greater, on average, for households that were initially multidimensionally poor. This suggests that emergency housing does not benefit all households in the same way, but rather the effects are magnified for the poorest households.

The analysis of the dimensional contributions to multidimensional poverty demonstrated how changes in the adjusted headcount measure are produced by the combined effects of moving households out of multidimensional poverty and the decline in the overall deprivation rate of each indicator. The Alkire-Foster method focuses attention on the situation of those households that experience an accumulation of deprivations.

It is important to consider several limitations when assessing the paper's results. First, the Alkire-Foster measure was defined only in terms of dimensions related to housing. If we had included dimensions such as work or education, on which housing has a more indirect effect, the treatment effects would likely have been far smaller. Second, we measured the programme's impact over a relatively short period of time, just one year. It is likely that over time the positive effects of the programme will dissipate, both due to the wear and tear on the basic dwelling and because the household member's enthusiasm over receiving a new dwelling is apt to decline over time. At the same time, interactions between wellbeing dimensions (e.g., between emotional health, interpersonal relations and sleep) could cause other effects to emerge and evolve over time. Third, although the programme produced a marked reduction in deprivations, an ample percentage of households continue to be deprived in multiple dimensions of wellbeing. Even after receiving the TECHO house, in over half of all households at least one bed is overcrowded; in one in five households the members do not have space to dress in private; and nearly half of all respondents experienced sadness, anxiety or depression during the previous month. All of these findings point to the essential need for public sector actions to improve access to quality public services and adequate housing for the most vulnerable households in Argentina.

Nonetheless, the evaluation results leave little doubt that TECHO's emergency housing programme produced a large reduction in the accumulation of deprivations in essential dimensions of human life. This positive effect may provide the impetus for these families to improve other dimensions of their lives or may be an intermediate step toward obtaining a permanent house, which is what many programme participants say that they long for the most.

References

- Alkire, S. and Foster, J. (2011). 'Counting and multidimensional poverty measurement', *Journal of Public Economics*, vol. 95(7–8), pp. 476–487.
- Alkire, S., Foster, J., Seth, S., Santos, M.E., Roche, J.M., and Ballon, P. (2015). *Multidimensional Poverty Measurement and Analysis*. Oxford: Oxford University Press.
- Alkire, S. and Santos, M.E. (2011). 'Acute multidimensional poverty: A new index for developing countries', United Nations Development Programme Human Development Research Paper 2010/1, New York.
- Anderson, M. L. (2008). 'Multiple inference and gender differences in the effects of early intervention: A re-evaluation of the Abecedarian, Perry Preschool, and Early Training Projects', *Journal of the American Statistical Association*, vol. 103(484), pp. 1481–1495.
- Angrist, J. and Pischke, J.S. (2009). *Mostly Harmless Econometrics: An Empiricist's Companion*. Princeton: Princeton University Press.
- Bartlett, S.N. (2002). 'The problem of children's injuries in low-income countries: A review', *Health Policy and Planning*, vol. 17(1), pp. 1–13.
- Benicio, M.H.D., Ferreira, M.U., Cardoso, M.R., Konno, S.C., and Monteiro, C.A. (2004). 'Wheezing conditions in early childhood: Prevalence and risk factors in San Paolo, Brazil', *Bulletin of the World Health Organization*, vol. 82, pp. 51–522.
- Bonnefoy, X. (2007). 'Inadequate housing and health: An overview', *Int. J. Environment and Pollution*, vol. 30(3), pp. 411–429.
- Brunekreef, B., Dockery, D.W., Speizer, F.E., Ware, J.H., Spengler, J.D., and Ferris, B.G. (1989). 'Home dampness and respiratory morbidity in children', *American Review of Respiratory Disease*, vol. 140, pp. 1363–1367.
- Buxton, O.M. and Marcelli, E. (2010). 'Short and long sleep are positively associated with obesity, diabetes, hypertension, and cardiovascular disease among adults in the United States', *Social Science and Medicine*, vol. 71, pp. 1027–1036.
- Cardinali, D. P., Espinola, G.S.S.R., Salvia, A., and Vigo, D.P.C.D.E. (2014). 'Sleep, slums and shelter: Impact of a slum-housing upgrading programme', *Proceedings of the Working Group in Bread and Brain: Education and Poverty*, pp. 1–14.

- Cattaneo, M.D., Galiani, S., Gertler, P.J., Martinez, S., and Titiunik, R. (2009). 'Housing, health and happiness', *American Economic Journal: Economic Policy*, vol. 1(1), pp. 75–105.
- Evans, G. (2003). 'The built environment and mental health', *Journal of Urban Health*, vol. 80(4), pp. 536–555.
- Fonseca W., Kirkwood B.R., Victoria C.G., Fuchs S.R., Flores, J.A., and Misago C. (1996). 'Risk factors for childhood pneumonia among the urban poor in Fortaleza, Brazil: A case-control study', *Bulletin World Health Organization*, vol. 74(2), pp. 199–208.
- Foster, J., Greer, J. and Thorbecke, E. (1984). 'A class of decomposable poverty measures', *Econometrica*, vol. 52(3).
- Galiani, S., Gertler, P., Undurraga, R., Cooper, R., Martinez, S., and Ross, A. (2016). 'Shelter from the storm: Upgrading housing infrastructure in Latin American slums', *Journal of Urban Economics*, vol. 96, pp. 166–194.
- Galiani, S. and Scharfrodsky, E. (2010) 'Property rights for the poor: Effects of land titling', *Journal of Public Economics*, vol. 94, pp. 700–729.
- Gove, W. R., Hughes, M., and Galle, O. R. (1979). 'Overcrowding in the home: An empirical investigation of its possible pathological consequences', *American Sociological Review*, vol. 44(1), pp. 59–80.
- Imbens, G. W. (2015). 'Matching methods in practice: Three examples', *Journal of Human Resources*, vol. 50(2), pp. 373–419.
- Krieger, J., and Higgins, D. L. (2002). 'Housing and health: Time again for public health action', *American Journal of Public Health*, vol. 92(5), pp. 758–768.
- Lehmann, E.L. and Romano, J. (2005). *Testing Statistical Hypotheses*. New York: Springer.
- Loschman, C., Parsons, C. and Siegel, M. (2015). 'Does shelter assistance reduce poverty in Afghanistan?', *World Development*, vol. 74, pp. 305–322.
- Mitchell, A., Macció, J. and Mariño Fages, D. (2016). 'El efecto de la vivienda de emergencia en el bienestar', *Anales de las Jornadas de la Asociación Argentina de Economía Política 2016*, Buenos Aires.
- Murtagh P., Cerqueiro C., Halac A., Avila M., Salomon H., and Weissenbacher M. (1993). 'Acute lower respiratory infection in Argentinian children: a 40-month clinical and epidemiological study', *Pediatric Pulmonology*, vol. 16(1), pp. 1–8.

- Newman, S. J. (2008). 'Does housing matter for poor families? A critical summary of research and issues still to be resolved', *Journal of Policy Analysis and Management*, vol. 27(4), pp. 895–925.
- Ortiz Flores, E. (2011). 'La producción social del hábitat' in (M. Arévalo et al., eds.), *El Camino Posible. Producción social de hábitat en América Latina*, Montevideo: Ediciones Trilce.
- Peat J.K., Dickerson J., and Li, J. (1998). 'Effects of damp and mould in the home on respiratory health: A review of the literature', *Allergy*, vol. 53, 120–128.
- Perneger, T. V. (1998) 'What's wrong with Bonferroni adjustments', *British Medical Journal*, vol. 316, pp. 1236–1238.
- Roche, J. M. (2013). 'Monitoring Progress in child poverty reduction: Methodological insights and illustration to the case study of Bangladesh', *Social Indicators Research*, vol. 112(2), pp. 363–390.
- Rojas, E. (2016). 'Housing policies and urban development. Lessons from the Latin American experience, 1960–2010', in (G.W. McCarthy, G. Ingram and S.A. Moody, eds.) *Land and the City*, Cambridge, MA: Lincoln Institute of Land Policy.
- Sen, A. (1985). 'Well-being, agency and freedom: The Dewey lectures 1984', *The Journal of Philosophy*, vol. 82(4), pp. 169–221.
- Simonelli, G., Leanza, Y., Boilard, A., Hyland, M., Augustinavicius, J. L., Cardinali, D. P., and Vigo, D. E. (2013). 'Sleep and quality of life in urban poverty: The effect of a slum housing upgrading programme', *Sleep*, vol. 36(11), pp. 1669–1676.
- Simonelli, G. Patel, S., Rodríguez-Espínola, S., Pérez-Chada, D., Salvia, A., Cardinali, D., and Vigo, D. (2015). 'The impact of home safety on sleep in a Latin American country', *Sleep Health*, vol. 1(2), pp. 98–103.
- Strine, T.W. and Chapman, D.P. (2005). 'Associations of frequent sleep insufficiency with health-related quality of life and health behaviors', *Sleep Medicine*, vol. 6, pp. 23–27.
- TECHO Argentina (2013). *Relevamiento de Asentamientos Informales 2013*. [Link](#).
- Thomson, H., Thomas, S., Sellstrom, E., and Petticrew, M. (2009). 'The health impacts of housing improvement: A systematic review of intervention studies from 1887 to 2007', *American Journal of Public Health*, vol.99 (Suppl. 3), pp. S681–S692.
- United Nations Development Programme (2015). *The Millennium Development Goals Report*. New York: United Nations. [Link](#). Accessed 10 September 2017.

- United Nations Office of the High Commissioner for Human Rights (2009). Fact Sheet No. 21. The Human Right to Adequate Housing. [Link](#). Accessed 10 September 2017.
- UN-Habitat (2016). *World Cities Report 2016. Urbanization and Development. Emerging Futures*. Nairobi: United Nations Human Settlements Programme.
- United Nations Sustainable Development Solutions Network (2018). Indicators and monitoring framework. Launching a data revolution for the Sustainable Development Goals. [Link](#). Accessed 5 April 2018.
- Wagner, R. F. (2011). 'La Producción Social del Hábitat en la ciudad injusta' in (M. Arévalo, et al., eds.), *El Camino Posible. Producción social de hábitat en América Latina*. Montevideo: Ediciones Trilce.
- White, H. and Barbu, A. (2006). *Impact Evaluation: The Experience of the Independent Evaluation Group of the World Bank*. Washington, DC: World Bank.
- Wolff, J. and De-Shalit, A. (2007). *Disadvantage*. Oxford: Oxford University Press.

Annex

Annex Table A1: Effect of treatment on indicators of the size and quality of the dwelling

Dimension/Indicator	Full sample		n
	Control Mean (Std. Dev.)	Treatment effect	
HOUSING SIZE			
Number of members	5.260 (2.432)	-0.033 (0.131)	546
M ² of the dwelling	29.560 (22.875)	6.980 *** (2.238)	538
M ² of the dwelling per person	6.525 (6.341)	1.745 *** (0.569)	538
Number of rooms	2.072 (1.169)	0.632 *** (0.137)	546
Persons per room	3.198 (1.963)	-0.979 *** (0.218)	546
Overcrowding (< 5 m ² per person)	0.477 (0.5)	-0.242 *** (0.066)	538
Overcrowding (> 2 persons per room)	0.638 (0.481)	-0.234 *** (0.047)	546
Overcrowding (> 3 persons per room)	0.365 (0.482)	-0.232 *** (0.052)	546
HOUSING QUALITY			
Share of rooms with good quality roof	0.885 (0.32)	0.008 (0.043)	539
Share of rooms with good quality walls	0.905 (0.294)	-0.025 (0.029)	538
Share of rooms with good quality floors	0.727 (0.446)	0.096 (0.074)	539
Share of rooms with problems with roof	0.763 (0.426)	-0.391 *** (0.054)	538
Share of rooms with problems with walls	0.737 (0.441)	-0.362 *** (0.066)	536
Share of rooms with dirt floor or seepage	0.389 (0.488)	-0.148 ** (0.069)	545
Share of rooms that flood when it rains	0.469 (0.5)	-0.245 *** (0.062)	534
Share of rooms very or quite humid	0.664 (0.473)	-0.384 *** (0.056)	538

Note: Data from TECHO Housing survey, 2014-15.